

1 Benedikt Szmrecsanyi and Alexandra Engel
2 **Register Variation in a Cognitive**
3 **(Socio)linguistics Perspective**

4 **Abstract:** Key questions in Cognitive (Socio)Linguistics include the following:
5 “How do language users acquire lectal competence, how is it stored mentally,
6 and how does it work in language production?” (Geeraerts, Kristiansen &
7 Peirsman 2010: 10). We aim to shed more light on the storage and production
8 component of this question. Specifically, we will explore the extent to which lan-
9 guage users have different probabilistic grammars for different situational varie-
10 ties of speech and writing (“registers”) – do our linguistic choice making pro-
11 cesses differ depending on whether we engage in e.g., informal conversation or
12 write blog entries? This issue is under-researched but loaded theoretically. Our
13 case study is about the dative alternation in English (John gave the president a
14 present versus John gave a present to the president). The methodology is usage-
15 based and relies on both corpus evidence (i.e., observation) and a rating task ex-
16 periment. We distinguish between four broad registers: informal spoken lan-
17 guage, formal spoken language, informal written language, and formal written
18 language. Analysis shows that different registers do indeed come with different
19 probabilistic grammars, which indicates that lectal/register differences play an
20 important role in cognitive categorization.

21 **Keywords:** register, probabilistic grammar, variation, dative alternation

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24 **1 Introduction**

25 This paper is about how the way people choose between “alternate ways of say-
26 ing ‘the same’ thing” (Labov 1972: 188) depends on the situation of spoken and
27 written language production, a.k.a. register. We stress that register variation as
28 one important manifestation of lectal variation (including the distinction

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29 between formal and informal text types) has been a key focus in Leuven-school
30 lectometry and cognitive sociolinguistics (see e.g., Speelman, Grondelaers and
31 Geeraerts 2006). Our point of departure then is that register variation is rampant
32 in human language (Ferguson 1983: 154), and that knowledge of how to use lan-
33 guage in particular situations is a key ingredient of language users' lectal
34 knowledge.

35 **2 State-of-the Art**

36 Previous research on register variation has primarily focused on the text frequen-
37 cies of particular linguistic features in particular registers: how often or rarely do
38 we find particular linguistic features, such as passive constructions, in particular
39 registers? The flagship method in this line of research is the Multi-Dimensional
40 (MD) approach developed by Douglas Biber (1988), which measures co-occur-
41 rence patterns of linguistic features. An alternative, variationist (in the spirit of,
42 e.g., Labov 1972; Grondelaers and Speelman 2007) way of approaching register
43 variation does not ask how frequently particular features are used in particular
44 registers, but the following question instead: when speakers can choose between
45 different ways of saying the same thing, what is the extent to which they draw on
46 different choice-making processes in different registers? Such probabilistic regis-
47 ter differences have received short shrift in the past. Variationist sociolinguists in
48 the tradition of Labov's work would have the methodological toolkit to investi-
49 gate these issues, but this community happens to be mostly interested in one par-
50 ticular register, vernacular speech as observable in sociolinguistic interviews
51 (but see e.g., D'Arcy and Tagliamonte 2015). Probabilistic effects also take center
52 stage in Probabilistic Grammar work à la Joan Bresnan and collaborators, but
53 again most of the extant work in this tradition is concerned with spoken language
54 (exceptions include e.g., Bresnan et al. 2007, Grafmiller 2014). In sum, the regis-
55 ter-sensitivity of probabilistic choice-making should be of central theoretical im-
56 portance to analysts working in experienced-based and usage-based paradigms,
57 but so far this sensitivity is ill-understood and in want of empirical investigation.

58 **3 Research Questions**

59 This paper conducts a pilot study for the sake of determining the extent to which
60 language users' probabilistic grammars may include knowledge about lectal

61 register differences. The goal is to investigate the degree to which language users'
 62 choice-making processes are different as a function of register. We thus aim to
 63 assess – via corpus analysis and rating task experiments – how the effect size and
 64 direction of language-internal constraints on variation interacts with register as
 65 a language-external parameter. Two more specific research questions guide our
 66 analysis in the present paper:

1. With regard to register distinctions – what are the relevant register-related dimensions of variability: formal versus informal (formality), or written versus spoken (medium)?
2. With regard to probabilistic constraints – what are the probabilistic constraints that tend to have particularly variable probabilistic effects across registers?

67 4 Methodology

68 The alternation we analyze by way of a case study is the dative alternation after
 69 the verb *give* in English. To encode dative relations, speakers and writers of Eng-
 70 lish may use two semantically roughly equivalent structural patterns: the ditransi-
 71 tive dative variant, as in (1a), and the prepositional dative variant, as in (1b):

- (1) a. *Several charities have different stances on whether or not you should give*
[homeless people]_{recipient} [money]_{theme} directly
 (The Independent, 10/01/2018)
- b. *Mm and I used to give [a lot of money]_{theme} [to homeless groups]_{recipient}*
 (BNC2014, SPHJ)

72 The dative alternation is one of the best-understood alternations in the grammar
 73 of English. One seminal study on the dative alternation in English is Bresnan et
 74 al. (2007), which explores factors that constrain language users' dative choices
 75 in American English (telephone) conversations. Bresnan et al. find that variation
 76 between the two dative options is constrained by about 10 predictors, including
 77 e.g., pronominality of the recipient/theme, discourse accessibility (pragmatics),
 78 and animacy of the recipient. If, for example, unlike in (1) the recipient is inani-
 79 mate, Bresnan et al. 's regression model predicts that the odds for the preposi-
 80 tional dative increase by a factor of about 4. This is the probabilistic effect that
 81 inanimate recipients have on dative choice in telephone conversations. But do
 82 inanimate recipients have the same effect in, say, formal speeches? What about

83 the other predictors? What is the extent to which language users have to adjust
 84 probabilistic decision-making when they switch from telephone conversations to
 85 other registers? These are the kinds of questions that we are interested in.

86 4.1 Corpus-based track

87 A corpus-based variationist analysis applying logistic regression to a richly an-
 88 notated dataset was carried out (see Szmrecsanyi 2019 for discussion). We chose
 89 four registers at the intersection between formality and mode. Focusing on British
 90 English, we selected the Spoken BNC2014 (~11.4 million words) for informal con-
 91 versations between friends and family members (Love et al. 2017); a corpus of
 92 Hansard transcriptions from House of Commons debates for formal spoken lan-
 93 guage (~59.4 million words) (Marx and Schuth 2010); the British English blogs
 94 part of the GloWbE corpus for informal written language (~148 million words)
 95 (Davies 2013); and a corpus of newspaper articles from *The Independent* (~113.5
 96 million words) representing formal written language (JSI Newsfeed corpus, Bušta
 97 et al.).

98 From these corpora, we automatically extracted tokens of the verb *give*,
 99 which were then manually checked for criteria of the variable context. Accord-
 100 ingly, tokens were filtered out which included only one constituent, mistaggings
 101 (e.g., *given* as preposition or adjective), non-canonical word order, clausal con-
 102 stituents, *give* as particle verb, fixed expressions, passivized or relativized con-
 103 structions, and constructions in which the *to*-phrase depended on the theme (as
 104 in *give the answer to that question*). For reasons of speaker/author contribution,
 105 direct quotes were also filtered out. From the remaining tokens, a balanced, ran-
 106 dom sample of 2,600 observations was created (i.e., 650 tokens per corpus, half
 107 of which were ditransitive dative constructions and the other half were preposi-
 108 tional dative constructions). The dataset was annotated for the following predic-
 109 tors: pronominality (pronominal vs. non-pronominal), animacy (animate vs. in-
 110 animate), definiteness (definite vs. indefinite), constituent length (in number of
 111 characters), complexity (simple vs. complex), and (head) frequency of both con-
 112 stituents as well as verb sense (transfer of concrete object, transfer of abstract
 113 object, communication sense).¹ Constituent length measures were combined into
 114 a single predictor, Weight Ratio, by dividing recipient length by theme length.

1 Collective nouns were annotated as ‘inanimate’. Complex constituents included restrictive postmodifications to the constituent’s head.

115 Numerical predictors were log-transformed and standardized to reduce multicol-
116 linearity.

117 A logistic mixed effects regression model was then fitted in *R* using the *lme4*
118 package (Bates et al. 2015). To test for the effect of register on the internal con-
119 straints, three interactions between Register and Weight Ratio, Recipient Defi-
120 niteness, and Theme Definiteness were included in the model in addition to the
121 main effects for all of the above predictors. All levels were set to the default levels
122 of the ditransitive dative. Random effects for recipient and theme head lemma as
123 well as for speaker identity account for idiosyncrasies. Model selection followed
124 a backward elimination process. The resulting model has a high *C* index of 0.97,
125 confirming outstanding model performance, and an accuracy of 91.4% (baseline
126 50%).

127 4.2 Experimental track

128 This corpus model was then tested against human rating performance. Are lan-
129 guage users sensitive to probabilistic patterns in the choice of dative variants?
130 More specifically, do we find similar patterns in a comparison between corpus-
131 based predictions and language users' intuitions about the probability of vari-
132 ants? To this end, we set up a rating task experiment in which we presented par-
133 ticipants with both variants in authentic corpus examples. Participants were
134 asked to give gradient ratings as to which variant they judge more likely in the
135 context. Previous research in this vein has found converging evidence between
136 corpus results and ratings (Bresnan and Ford 2010; see Klavan and Divjak 2016
137 for a review). In a seminal study on the dative alternation, Bresnan and Ford
138 (2010) asked American English and Australian English speakers to distribute 100
139 points across both variants and found variety-specific probabilistic effects corre-
140 sponding to the patterns found in the corpus. Do we find similar patterns when
141 we examine register-specific knowledge?

142 Based on the corpus results, spoken informal and spoken formal registers
143 were chosen to provide the items for the present experiment (see Section 5.1). In
144 total, 32 corpus excerpts were selected, 16 for each register. Per register, six items
145 contained dative constructions and ten filler items contained either lexical, reg-
146 ister-specific choices (four items) or the choice between the relativizers *which* and
147 *that* (six items). These fillers were included to distract from the target construc-
148 tions. Target items came from six probability bins across the whole probability
149 range (probability of 0.06-0.99 for the prepositional dative). In order to control
150 for possible confounding variables, the target items were restricted to those items
151 that included simple, non-pronominal constituents and a definite recipient. As a

152 result, three items per register were not included in the dataset and were thus
 153 unseen by the corpus model. Per register, target items were counterbalanced for
 154 dative variant, theme definiteness and whether they were seen or unseen by the
 155 corpus model. The full item set will be published as part of Engel et al. (in prepa-
 156 ration).

157 Two lists were created by varying the presentation side of the original variant.
 158 All items of one register were presented after one another followed by all items of
 159 the other register. Per list, there were two versions to account for possible order
 160 effects: version A began with the formal spoken items, version B with the informal
 161 spoken items. Eight simple comprehension questions were included to ensure
 162 that participants read the excerpts carefully.

163 The rating task was implemented in an online survey using Qualtrics. Partic-
 164 ipants were sampled via Qualtrics Research Services. One hundred British Eng-
 165 lish monolingual native speakers (50 female, 50 male; mean age: 55 years old;
 166 age range: 19-78) took part in the study and gave informed written consent before
 167 completing the survey. Participants were asked to indicate their ratings by means
 168 of a slider bar. Mean duration was 26 minutes.

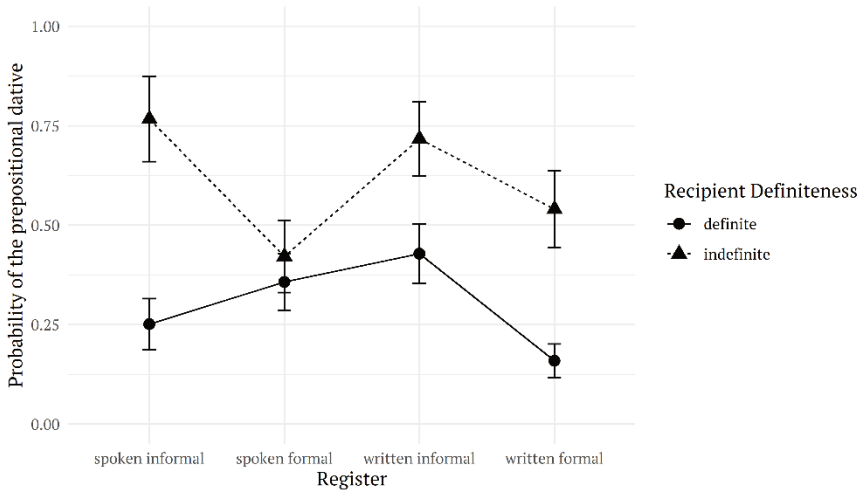
169 For the analysis, eleven participants were excluded either due to low accu-
 170 racy ($< 75\%$ correct answers) in response to the comprehension questions or due
 171 to excessive time spent on the survey (> 40 minutes; based on interquartile
 172 range). Ratings were standardized and entered as dependent variable in a linear
 173 mixed effects regression model with Predicted Corpus Probability and Weight Ra-
 174 tio, and an interaction between Register and Theme Definiteness as explanatory
 175 variables. In addition, a random effect of participant with Predicted Corpus Prob-
 176 ability in the slope was added to account for participant-specific variability.

177 5 Results

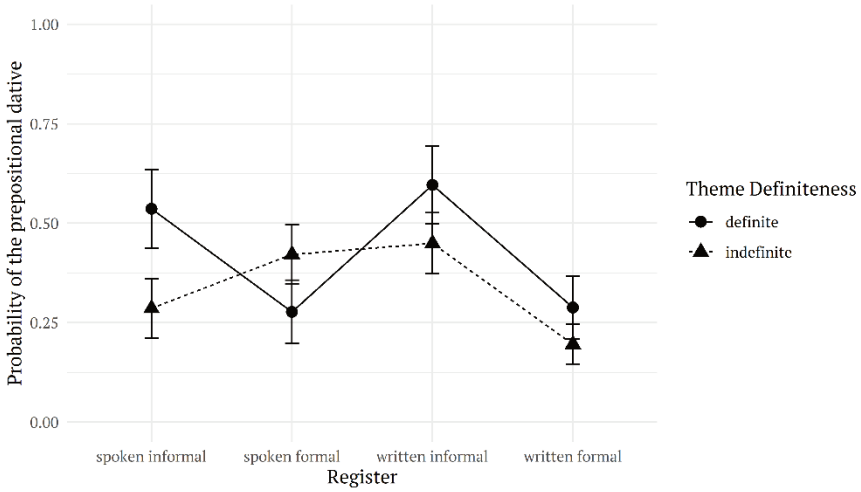
178 5.1 Corpus-based track

179 The corpus model indicates that the prepositional dative becomes more likely
 180 when the recipient is inanimate ($\beta = 0.95, p < .001$), indefinite ($\beta = 2.28, p < .001$),
 181 non-pronominal ($\beta = 2.18, p < .001$), complex ($\beta = 0.68, p = .02$), longer than the
 182 theme ($\beta = 1.84, p < .001$) and when the theme is simple ($\beta = 2.19, p < .001$), definite
 183 ($\beta = 1.06, p = .005$), and pronominal ($\beta = 2.18, p = .005$). Main effects for Verb
 184 Sense and Register indicate that the prepositional dative becomes more likely
 185 when *give* has a communication sense ($\beta = 1.37, p < .001$) or a transfer sense ($\beta =$
 186 $0.73, p = .011$), and that the probability of a prepositional dative with *give* is higher

187 in spoken formal ($\beta = 1.06, p = .005$) and written informal ($\beta = 0.96, p = .008$)
 188 registers compared to the spoken informal register. Moreover, there are interactions
 189 between Register and Recipient Definiteness (Figure 1) and between Register
 190 and Theme Definiteness (Figure 2). These interactions show that the effect size
 191 of Recipient Definiteness is modulated in the spoken formal register compared to
 192 the spoken informal register ($\beta = -2.02, p = .003$) and that the direction of the effect
 193 of Theme Definiteness is reversed in the spoken formal register ($\beta = -1.7, p =$
 194 $.001$). The random effect for theme head lemma significantly contributes to explaining
 195 the variation ($\sigma^2 = 3.27$).



196
 197 **Fig. 1:** Interaction effect between register and recipient definiteness in corpus model. The prob-
 198 ability of the prepositional dative (y-axis) increases when the recipient is indefinite across all
 199 registers (x-axis), but the magnitude of the effect is modulated in the spoken formal register.
 200 Error bars represent standard errors.

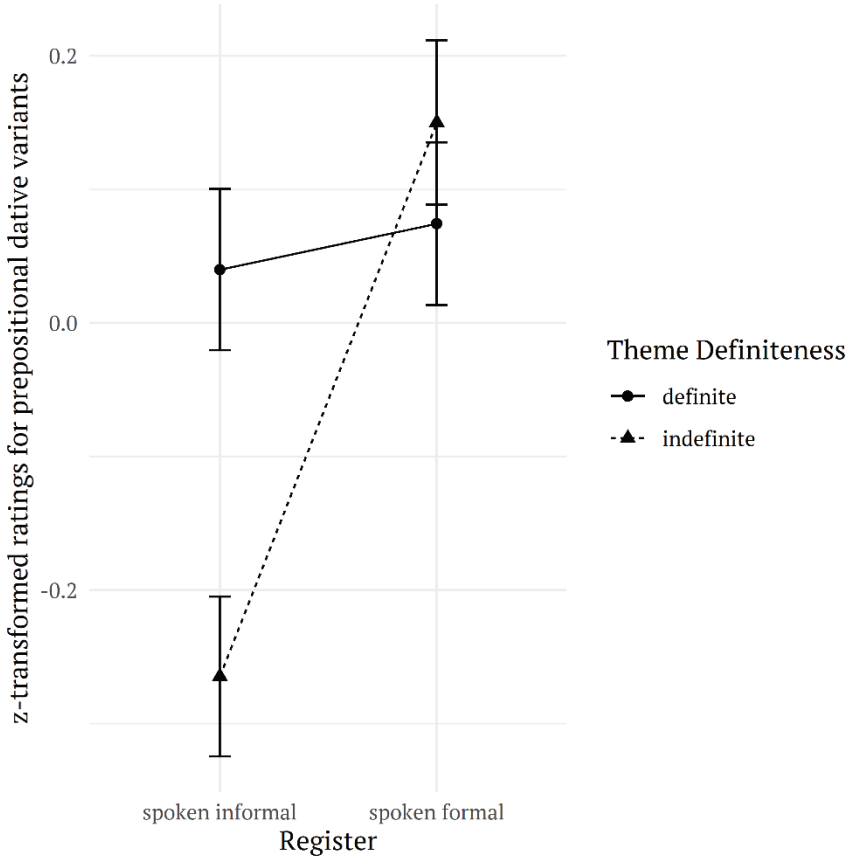


201

202 **Fig. 2:** Interaction effect between register and theme definiteness in corpus model. The proba-
 203 bility of the prepositional dative (y-axis) increases when the theme is definite except in the
 204 spoken formal register where the direction of the effect is reversed. Error bars represent stand-
 205 ard errors.

206 5.2 Experimental track

207 There is a main effect for Predicted Corpus Probability ($\beta = 0.3, p < .001$), indicat-
 208 ing that participants gave higher ratings for the prepositional dative as the pre-
 209 dicted probability for the prepositional dative in the corpus model increases. In
 210 addition, there is an interaction between Register and Theme Definiteness ($\beta =$
 211 $0.39, p = .001$; see Figure 3). Participants gave higher ratings for the prepositional
 212 dative in spoken formal items with indefinite themes. These results show that the
 213 corpus model and participants' ratings pattern together. There were also main
 214 effects for Weight Ratio ($\beta = -0.19, p < .001$), and Theme Definiteness ($\beta = -0.33, p$
 215 $< .001$). Note that with a conditional R^2 of 0.11 and a marginal R^2 of 0.09, the
 216 model leaves a large part of the variance unexplained, which might be due to
 217 individual variation in the extent to which participants made use of the rating
 218 scale.



219

220 **Fig. 3:** Interaction between register and theme definiteness in participants' responses. Partici-
 221 pants gave higher ratings for the prepositional dative in spoken informal items when the
 222 theme was definite in contrast to higher ratings for the prepositional dative in spoken formal
 223 items with indefinite themes. Ratings are expressed as z-scores.

224 Results for the filler items show an interaction between Register and Filler Type
 225 ($\beta = -0.39$; $p = 0.002$). In the spoken formal register, participants had stronger
 226 preferences for the formal variant in lexical items compared to relativizer items.
 227 Additionally, a main effect for Register ($\beta = 0.9$; $p < .001$) indicates that formal
 228 variants overall received higher ratings in the formal register. These results con-
 229 firm that participants made register-specific judgments.

230

6 Discussion

231 Analysis shows that the main effects are in line with “harmonic alignment” (Bres-
 232 nan et al. 2007; Bresnan and Hay 2008; Theijssen et al. 2013; Röthlisberger,
 233 Grafmiller and Szmrecsanyi 2017 and/or “Easy First” effects (MacDonald 2013):
 234 early constituents tend to be simple, short, animate, and definite. But what about
 235 interactions with register? According to corpus data, register interacts with the
 236 effect of definiteness:

- In all registers, the prepositional dative is more likely when the recipient is indefinite, but the largest definiteness effect can be observed in the spoken informal register, while we find the smallest effect in the spoken formal register.
- As to the theme, the prepositional dative is overall more likely when the theme is definite, but the direction of the effect is reversed in the spoken formal register. As to effect sizes, we observe the largest effect in the spoken informal register, and the smallest effects in both formal registers.

237 The experimental analysis partially confirms the existence of these differences:
 238 in the rating data as well, we see an interaction between register and theme defi-
 239 niteness. Register-specific effects are subtle, but subjects still seem to be sensitive
 240 to such subtle effects. That said, there is a great deal of individual variation.

241 Why does definiteness interact with register? Supplementary analysis
 242 demonstrates that indefinite recipients are particularly frequent in the ditransi-
 243 tive dative in the spoken formal register. Assuming that definite referents are
 244 more accessible than indefinite ones (Gundel, Hedberg and Zacharski 2001), we
 245 may argue that in spontaneous conversation, definite referents are placed first
 246 because they are easier to access and to process. We also find more indefinite
 247 themes in the spoken formal register overall, and with the prepositional dative in
 248 particular; it thus seems that in general, more indefinite referents are used in par-
 249 liamentary debates compared to informal conversations. This might be explained
 250 by the high frequency of definite pronouns in informal conversations, as opposed
 251 to the higher frequency of nouns in more informational registers (Biber et al.
 252 1999: 235).

253 We now move on to a discussion of the wider significance of these results.
 254 Our findings have implications and relevance for theory formation. Our research
 255 is ultimately concerned with the nature and scope of linguistic knowledge, and
 256 with the interaction of this knowledge with socioculture (for register conventions

257 are social in nature). Generally speaking, variationist sociolinguists believe that
 258 “internal constraints [...] are normally independent of social and stylistic factors”
 259 (Labov 2010: 265), and it is of course this independence that our findings call into
 260 question. Given that definiteness as a probabilistic constraint has demonstrably
 261 different effect sizes (and sometimes even effect directions) across registers,
 262 Guy’s Grammatical Difference Hypothesis (Guy 2015), according to which having
 263 different constraints means having different grammars, would arguably warrant
 264 us to conclude that language users have a number of different register-specific
 265 grammars, akin to situations of diglossia or bilingualism. So, coming back to the
 266 cognitive sociolinguistics research question spelled out in the abstract – How is
 267 lectal competence stored mentally, and how does it work in language produc-
 268 tion? (Geeraerts, Kristiansen and Peirsman 2010: 10) – our analysis would seem
 269 to suggest that competence about register differences is maybe more crucial and,
 270 in fact, richer than previously assumed by many: if different register come with
 271 different (probabilistic) grammars as we have shown, then register competence
 272 is no different from multilingual or multidialectal competence.

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